

STSCI 6340:

Lecture 8

Date: 03/04

Announcements:

- Homework: Thursday April 8 (4pm)
return: next Monday (6pm) 12th.
 - Month or so to read the paper.
one-on-one meetings March 22.
 - Presentations will start on the week
of April 12. and end week of May 6.
 - Final week: no classes
 - Deadline for projects: Friday May 14.
(midnight).
 - Project: write up to 2 pages in
Nurips latex format.
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- Approximate message passing:

$$A \in \mathbb{R}^{n \times n} \text{ symmetric.}$$

$$\text{Initialize: } x^0 \in \mathbb{R}^n, \quad x^{-1} = 0 \quad f_0 = 0$$

$$\text{Iteration: } x^{t+1} = A f_t(x^t) - b_t \cdot f_{t-1}(x^{t-1}).$$

$f_t: \mathbb{R} \rightarrow \mathbb{R}$ Lipschitz continuous, differentiable.

$$f_t(x^t) = \left(f_t(x_i^t) \right)_{i=1}^n$$

$$b_t = \frac{1}{n} \sum_{i=1}^n f_t'(x_i^t) \left(\begin{array}{l} \text{remark:} \\ \text{Can replace } b_t \text{ by} \\ \mathbb{E}[f_t'(\mu_t x_0 + \sigma_t z)] \end{array} \right)$$

- Non linear power iteration:

$$\left. \begin{array}{l} x^{t+1} = A \hat{x}^t \\ \hat{x}^{t+1} = \frac{x^{t+1}}{\|x^{t+1}\|_2} \end{array} \right\} \quad f(x) = \frac{x}{\|x\|_2}$$

Feature of the algorithm: presence of

the "memory" term: $- b_t f_{t-1}(x^{t-1})$.

will ensure the iterates are Gaussian under distributional assumptions on A .

Assumptions the matrix A :

$$Y = \sqrt{\frac{\lambda}{n}} x_0 x_0^T + W \quad w_{ij} \sim N(0, 1)$$

$$A = \frac{1}{n} Y$$

$$w_{ij} = w_{ji} \\ w_{ii} \sim N(0, 2)$$

More generally: $w_{ij} = w_{ji}$ $\mathbb{E}[w_{ij}] = 0$
 $\mathbb{E}[w_{ij}^2] = 1$

w_{ij} is sub gaussian

$$\left(\mathbb{E}[e^{\lambda w_{ij}}] \leq e^{\frac{\lambda^2 \sigma^2}{2}} \forall \lambda \right)$$

• Exactly characterize the law of the iterates in the limit $n \rightarrow \infty$.

• Theorem: If A is as above,

The signal x_0 and X_{init}^0 have an empirical distribution of entries which converges weakly to a distribution P_0 on \mathbb{R} with bounded second moment.

Let $\hat{X}^t = f_t(x^t)$; and assume

$$\frac{1}{n} \|\hat{X}_{\text{init}}^0\|_2^2 \rightarrow \rho, \quad \frac{1}{n} \hat{X}_{\text{init}}^0 x_0 \rightarrow \mathcal{P}$$

Then for every pseudo-Lipschitz function $\psi: \mathbb{R}^2 \rightarrow \mathbb{R}$ we have

$$\forall t \geq 0 \quad \frac{1}{n} \sum_{i=1}^n \psi(x_i^t, x_{0i}) \rightarrow \mathbb{E}[\psi(\mu_t x_0 + \sigma_t z, x_0)]$$

where $x_0 \sim P_0$, $z \sim N(0, 1)$ independent.

- (μ_t, σ_t) satisfy the state evolution iteration.

$$\begin{cases} \mu_{t+1} = \Gamma \mathbb{E} \left[x_0 f_t'(\mu_t x_0 + \sigma_t z) \right] \\ \sigma_{t+1}^2 = \mathbb{E} \left[f_t^2(\mu_t x_0 + \sigma_t z) \right] \\ \mu_0 = \Gamma x \quad , \quad \sigma_0^2 = \rho \end{cases}$$

Definitions:

- $x^{(n)} \in \mathbb{R}^n$: The empirical distribution of the entries of $x^{(n)}$ is

$$m_n = \frac{1}{n} \sum_{i=1}^n \delta_{x_i^{(n)}}.$$

→ A sequence of probability measures $(m_n)_n$.

$m_n \xrightarrow{\text{weakly}} \rho_0$ has finite second moment.

- ψ is a pseudo-Lipschitz function if

$$|\psi(x) - \psi(y)| \leq L(1 + \|x\|_2 + \|y\|_2) \|x - y\|_2.$$

An example is $\psi(x) = x^2$.

- Choice of f_t : Suppose you're interested in estimating x_0 (signal).

$\left. \begin{array}{l} \text{maximize overlap between } x^* \text{ and } x_0 \\ \text{minimize MSE} \end{array} \right\}$

observed

- Suppose we $\vee x_0$ through the Gaussian additive channel $y = \mu_t x_0 + \sigma_t z$.

Best estimator is posterior mean:

$$\mathbb{E}[x_0 | \tilde{y}] = \frac{\int x e^{\frac{\mu_t}{\sigma_t^2} \tilde{y} x - \frac{\mu_t^2}{2\sigma_t^2} x^2} dP_0(x)}{\int e^{\frac{\mu_t}{\sigma_t^2} \tilde{y} x - \frac{\mu_t^2}{2\sigma_t^2} x^2} dP_0(x)}$$

let

$$g(y, \mu) = \frac{\int x e^{\frac{y}{\sigma^2} \mu x - \frac{\mu^2}{2\sigma^2} x^2} dP_0(x)}{\int e^{\frac{y}{\sigma^2} \mu x - \frac{\mu^2}{2\sigma^2} x^2} dP_0(x)}$$

Then $\mathbb{E}[x_0 | \tilde{y}] = g\left(\frac{\tilde{y}}{\sigma_t}, \left(\frac{\mu_t}{\sigma_t}\right)^2\right)$

Let

$$f_t(y) = g\left(\frac{y}{\sigma_t}, \left(\frac{\mu_t}{\sigma_t}\right)^2\right) \quad \forall y \in \mathbb{R}$$

our estimator is $\hat{x}^t = f_t(x^t)$
 x^t : t^{th} iterate of AMP.

$\hat{x}^t \cdot x_0$, $\|\hat{x}^t\|_2^2$? let $\tilde{y} = \mu_t x_0 + \sigma_t z$

$$\begin{aligned} \mu_{t+1} &= \sqrt{\lambda} \mathbb{E} \left[x_0 f_t(\mu_t x_0 + \sigma_t z) \right] \\ &= \sqrt{\lambda} \mathbb{E} \left[x_0 g\left(\frac{\tilde{y}}{\sigma_t}, \left(\frac{\mu_t}{\sigma_t}\right)^2\right) \right] \\ &= \sqrt{\lambda} \mathbb{E} \left[x_0 \cdot \mathbb{E} [x_0 | \tilde{y}] \right] \\ &= \sqrt{\lambda} \mathbb{E} \left[\mathbb{E} [x_0 | \tilde{y}]^2 \right] \quad \leftarrow \text{tower property} \\ &= \sqrt{\lambda} \mathbb{E} \left[f_t(\mu_t x_0 + \sigma_t z)^2 \right] \\ &= \sqrt{\lambda} \sigma_{t+1}^2 \end{aligned}$$

Define $q_t = \frac{1}{\lambda} \left(\frac{\mu_t}{\sigma_t} \right)^2$.

$$q_{t+1} = \frac{1}{\lambda} \frac{\mu_{t+1}}{\sigma_{t+1}^2} \cdot \mu_{t+1}$$

$$= \frac{1}{\lambda} \sqrt{\lambda} \cdot \sqrt{\lambda} \cdot \mathbb{E} \left[f_t(\mu_t x_0, \sigma_t z)^2 \right]$$

$$= \mathbb{E} \left[g \left(\frac{\mu_t x_0 + \sigma_t z}{\sigma_t}, \left(\frac{\mu_t}{\sigma_t} \right)^2 \right)^2 \right]$$

$$= \mathbb{E} \left[g \left(\frac{\mu_t}{\sigma_t} x_0 + z, \left(\frac{\mu_t}{\sigma_t} \right)^2 \right)^2 \right]$$

$$= \mathbb{E} \left[g \left(\sqrt{\lambda q_t} x_0 + z, \lambda q_t \right)^2 \right]$$

$$\psi(r) = \mathbb{E} \log \int e^{\sqrt{r} z x + r x x_0 - \frac{r}{2} x^2} dP_r(x)$$

$$y = \sqrt{r} x_0 + z$$

$$\psi'(r) = \frac{1}{2} \mathbb{E} \left[\mathbb{E} [x_0 | y]^2 \right]$$

Therefore $\boxed{q_{t+1} = 2 \psi'(\lambda q_t)}$

Let's understand of $\hat{x}_t^+ \cdot x_0$ and $\|\hat{x}_t^+\|_2^2$.

$$\hat{x}_t^+ \cdot x_0 = \frac{1}{n} \sum_{i=1}^n f_t(x_i^+) \cdot x_{0i}$$

Now let $\psi(x, x_0) = f_t(x) \cdot x_0$.

Apply the theorem:

$$\begin{aligned} \hat{x}_t^+ \cdot x_0 &\longrightarrow \mathbb{E} \left[x_0 f_t(\mu_t x_0 + \sigma_t z) \right] \\ &= \frac{\mu_{t+1}}{\sqrt{\lambda}} = q_{t+1} \end{aligned}$$

Now $\frac{1}{n} \|\hat{x}^t\|_2^2$.

Apply the theorem to $\psi(x, x_0) = f_t(x)^2$.

$$\Rightarrow \frac{1}{n} \|\hat{x}^t\|_2^2 \longrightarrow \mathbb{E} [f_t(\mu_t x_0 + \sigma_t \beta)^2]$$

$$= q_{t+1}$$

The normalized overlap:

$$\frac{\langle \hat{x}^t, x_0 \rangle}{\|\hat{x}^t\|_2 \cdot \|x_0\|_2} \longrightarrow \frac{q_{t+1}}{\sqrt{q_{t+1}}} = \sqrt{q_{t+1}}$$

q_t will converge to a fixed point of the function $F(q) = \lambda \psi(\lambda q)$.

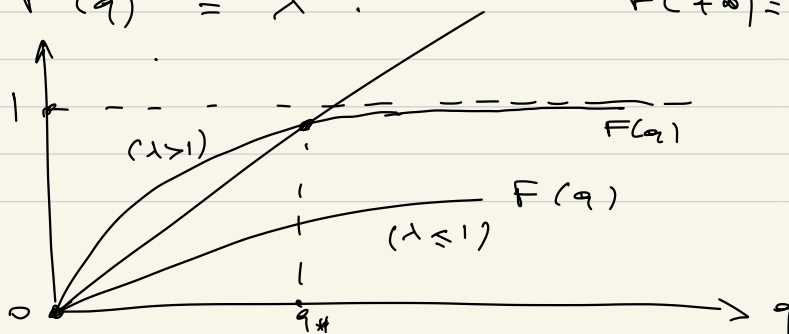
$$F(0) = \mathbb{E}[x_0]^2 = 0.$$

$$\psi''(\Gamma) = \frac{1}{2} \mathbb{E} [\text{cov}(x|y)^2]$$

$$\Gamma = 0 \quad ; \quad \psi''(0) = \frac{1}{2} \mathbb{E}[x^2] = \frac{1}{2}.$$

$$\Rightarrow F'(q) = \lambda.$$

$$F(+\infty) = \mathbb{E}[x_0^2] = 1$$

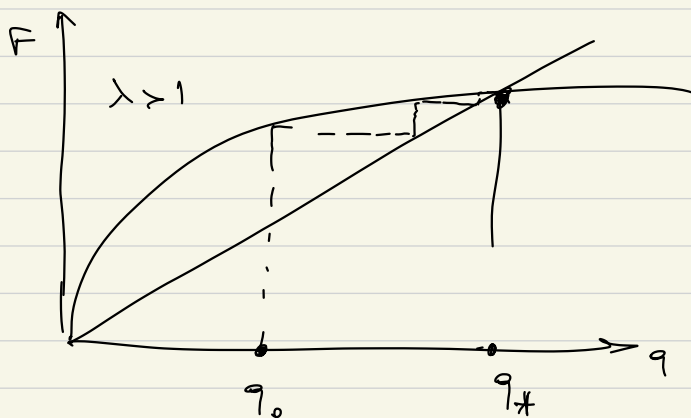


$\lambda \leq 1$: only one fixed point $q_* = 0$

$\lambda > 1$: Two fixed points 0 and $q_* > 0$.

AMP fails at estimating x_0 if $\lambda \leq 1$.

on the other hand if $\lambda > 1$ then it depends on initialization.



Let's say $y_i = \mathbb{E} x_{0i} + \varepsilon_i$ $\varepsilon > 0$

Initialize $\hat{x}_i^0 = \mathbb{E}[x_{0i} | y_i]$. $\forall 1 \leq i \leq n$.

$$q_0 = \mathbb{E} \psi'(\varepsilon) > 0.$$

overlap under

AMP $\rightarrow q_* = \text{Best overlap possible}$.

First remark: if $\lambda_c < 1$ then

AMP fails in $[\lambda_c, 1]$

but it is information-theoretically possible to estimate x_0 non trivially.

• Second remark: Initialize using first eigvec

of γ . u : first eigvec. Moutanari
- Venkateshramanian.

$$\text{Let } x^0 = \sqrt{n(\lambda^2 - 1)} \cdot u. \quad 2019.$$

$$q_0 = 1 - \frac{1}{\lambda} > 0 \quad \text{if } \lambda > 1.$$